Payments for care at private for-profit and private not-for-profit hospitals: a systematic review and meta-analysis

P.J. Devereaux, Diane Heels-Ansdell, Christina Lacchetti, Ted Haines, Karen E.A. Burns, Deborah J. Cook, Nikila Ravindran, S.D. Walter, Heather McDonald, Samuel B. Stone, Rakesh Patel, Mohit Bhandari, Holger J. Schünemann, Peter T.-L. Choi, Ahmed M. Bayoumi, John N. Lavis, Terrence Sullivan, Greg Stoddart, Gordon H. Guyatt

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Abstract

- **Background:** It has been shown that patients cared for at private for-profit hospitals have higher risk-adjusted mortality rates than those cared for at private not-for-profit hospitals. Uncertainty remains, however, about the economic implications of these forms of health care delivery. Since some policy-makers might still consider for-profit health care if expenditure savings were sufficiently large, we undertook a systematic review and meta-analysis to compare payments for care at private forprofit and private not-for-profit hospitals.
- **Methods:** We used 6 search strategies to identify published and unpublished observational studies that directly compared the payments for care at private for-profit and private not-forprofit hospitals. We masked the study results before teams of 2 reviewers independently evaluated the eligibility of all studies. We confirmed data or obtained additional data from all but 1 author. For each study, we calculated the payments for care at private for-profit hospitals relative to private notfor-profit hospitals and pooled the results using a random effects model.
- **Results:** Eight observational studies, involving more than 350 000 patients altogether and a median of 324 hospitals each, fulfilled our eligibility criteria. In 5 of 6 studies showing higher payments for care at private for-profit hospitals, the difference was statistically significant; in 1 of 2 studies showing higher payments for care at private not-for-profit hospitals, the difference was statistically significant. The pooled estimate demonstrated that private for-profit hospitals were associated with higher payments for care (relative payments for care 1.19,

95% confidence interval 1.07–1.33, p = 0.001).

Interpretation: Private for-profit hospitals result in higher payments for care than private not-for-profit hospitals. Evidence strongly supports a policy of not-for-profit health care delivery at the hospital level.

higher payments fo 95% confidence inter Interpretation: Private ments for care than strongly supports a p at the hospital level. *CMAJ* 2004;170(12):1817-24 S eparating issues of funding (i.e., who pays for health care) and delivery (i.e., who owns and administers the institutions providing care) helps to inform debates about health care systems. Funding for health care can come through private sources, primarily administered through insurance companies, or through public payment, by governments using tax dollars. Care can be delivered at private for-profit institutions that are owned by investors; private not-for-profit institutions that are owned by communities, religious organizations or philanthropic groups; or public health care institutions owned and administered by the government.

Canadian hospitals are publicly funded. In terms of delivery, although they are commonly referred to as public institutions, Canadian hospitals are almost all owned and operated by private not-for-profit organizations.¹ Canadian policy-makers continue to consider an expansion of private for-profit health care delivery, including private for-profit hospitals.¹

We have previously demonstrated higher risk-adjusted death rates among patients receiving care at private forprofit hospitals than among patients at private not-forprofit hospitals in a comprehensive systematic review.² Uncertainty remains, however, about the economic implications of these forms of health care delivery. Studies evaluating the economics of health care delivery usually evaluate costs, charges or payments for care.³ From the perspective of a service provider, costs represent how much the provider paid to provide care, charges represent how much the provider billed the payer, and payments represent how much the provider received for the care received. In the context of publicly funded health care, the central policy question is how much government will pay for care delivered by private for-profit versus private notfor-profit providers. We therefore undertook a systematic review and meta-analysis to address the following question: is there a difference in payments for patient care received at private for-profit compared with private not-forprofit hospitals?

Methods

We included published and unpublished observational studies and randomized controlled trials that directly compared payments for care at private for-profit and private not-for-profit hospitals. Because we required an estimate of variance to determine the precision of the estimate of the relative payments and to combine studies in our meta-analysis, we excluded from the quantitative meta-analysis studies that did not report (or whose authors could not provide) an estimate of variance for the payments.

Strategies to identify studies included an electronic search of 11 bibliographical databases; consultation with experts; review of our own files; review of reference lists from articles fulfilling our eligibility criteria; use of the "see related articles" feature in PubMed (in June 2003) for publications fulfilling our eligibility criteria; and use of SciSearch (in June 2003) for publications fulfilling our eligibility criteria.

A medical librarian used all the studies of which we were initially aware to identify medical subject-heading terms and key words for the search. In each database, the librarian iteratively refined the search strategy through testing several search terms and incorporating new search terms as new relevant citations were identified. The search included the following databases: EMBASE (1980–2001), MEDLINE (1966–2001), HealthSTAR (1975– 2001), CINAHL (1982–2001), BIOETHICSLINE (1973–2000), Wilson Business Abstracts (1997–2001), EconLit (1969–2001), Cochrane Library (2001, issue 3), Dissertation Abstracts Ondisc (1861–2001), ABI/Inform (1970–2001) and NTIS (National Technical Information Service) (1964–2002). The database search strategies are described in online Appendix 1 (www.cmaj.ca/cgi /content/full/170/12/1817/DC1).

Our 6 search strategies identified 7535 unique citations. Ten teams of 2 people independently screened the titles and abstracts of each citation and identified all citations that might contain a comparison of interest. This process yielded 788 full-text publications, identified by either screener, which we selected for full review (Fig. 1).

We masked the results by blacking them out in the tables and text of all publications selected for full review. To determine eligibility, 10 teams of 2 reviewers independently evaluated masked articles that they had not assessed during the screening process. The kappa value for agreement on article eligibility was 0.75. The consensus process to resolve disagreements required reviewers to discuss the reasoning for their decisions; in all cases, a realization of error by 1 reviewer completed the process. When both reviewers were uncertain as to whether a study was eligible, we contacted the author to clarify information.

Two reviewers independently abstracted the following data from all studies meeting eligibility criteria: sampling method, source of data, case mix, type of hospitals evaluated (e.g., general acute care, psychiatric), dates when data collection was initiated and completed, duration of patient follow-up, number of hospitals and patients evaluated, patient source of payment (e.g., public, private insurance) and potential confounders adjusted for in the analyses. Reviewers resolved disagreements by consensus using the process described above. Our overall agreement was 89% for the data abstraction. We attempted to contact the authors of all eligible studies to obtain or confirm data.

Before undertaking this systematic review we considered studies methodologically strong if they adjusted for the following factors: age, sex, ethnicity, income, education, primary diagnosis (case mix), comorbid conditions, severity of illness, market competition (the concentration of hospitals in a region), patient source of payment and hospital teaching status. We also considered analyses to be overadjusted if investigators adjusted for a variable that was under the control of hospital administrators, could vary by profit status and could possibly affect payments for care. For example, we considered an analysis overadjusted if it adjusted for staffing levels or the skill level of the hospital staff. Our quality assessment of studies included whether the study appropriately adjusted for any of the factors listed above and avoided overadjustment.

Before the analysis, we specified several hypotheses to explain variability in the direction and magnitude of effect across studies. We hypothesized that the effect size may differ depending on whether the hospitals evaluated were specialty or general hospitals; whether the payments for care were per discharge or per day; whether the payments for care were related to the hospital stay or included a period of time after hospital discharge; whether the analysis adjusted for potential confounders or was unadjusted; whether the patient source of payment was public or mixed (public and private); whether the patient population was adult or pedi-

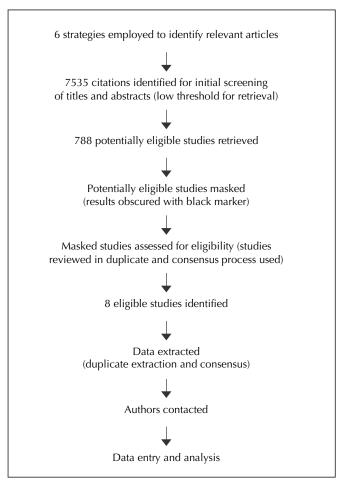


Fig. 1: Methodological steps of our systematic review.

atric; and whether the data collection occurred before 1984 (when US Medicare switched from reimbursement based on the costs of care to reimbursement based on a patient's diagnosis).

For each study we computed the relative payments for care at private for-profit hospitals relative to private not-for-profit hospitals (see online Appendix 2 at www.cmaj.ca/cgi/content/full/170/12 /1817/DC2). If a study reported 2 or more analyses with variance data, we included the adjusted analysis over an unadjusted analysis, and the analysis based on payments per discharge over that based on payments per day. We pooled these relative payments for care using a random effects model⁴ by weighting the natural logarithm of the relative payments by the inverse of their variances. Relative to the fixed effects model, the random effects model allows for between-study variation in the effect measure in addition to withinstudy variation. The random effects model generally results in wider confidence intervals than the fixed effects model. Although the random effects model still gives greater weight to studies with smaller variance than to studies with larger variance, the relative weight assigned to large studies is reduced compared with the weight assigned to those studies in the fixed effects model.

We calculated an I² as a measure of heterogeneity for the main analysis. An I² value represents the percentage of total variation across studies that is caused by heterogeneity rather than by chance. We considered a low I² value as 25% or lower and a high I² value as 75% or higher.⁵ We conducted a visual examination of funnel plots for evidence of publication bias.⁶

The Hamilton Health Sciences Research Ethics Board in Hamilton, Ontario approved the study protocol.

Results

We identified 8 publications of observational studies that met our eligibility criteria.⁷⁻¹⁴ We also identified 10 publications that we believed might be eligible but which required additional data from the authors. We were able to contact 9 of the 10 authors and to confirm that 6 of these studies had no measure of variance, 3 grouped private notfor-profit and public not-for-profit hospitals together and the authors either no longer had the data or could not rerun the analyses excluding the public hospitals, and 1 study had data on charges for care and did not have data on payments for care; these 10 studies were excluded from our systematic review (Table 1).^{15–24} All 10 excluded studies found higher payments or charges for care at private forprofit hospitals than at not-for-profit hospitals; in 6 of the 10, the differences were statistically significant.^{15–17,21,22,24}

Table 2 presents the study characteristics and Table 3 the study methodology of the 8 observational studies included in our systematic review. We obtained or confirmed data with the investigators for 7 studies; the sole author of the remaining study had died.⁸ All studies were conducted in the United States and included data from 1980 until Dec. 31, 1994. The 8 studies included over 350 000 patients and assessed a median of 324 hospitals per study.

Our quality assessment of studies revealed that 6 of the 8 studies appropriately adjusted or matched cases for many important determinants of payment for care (e.g., case mix). For 2 studies the authors were unable to provide an estimate of variance for their adjusted analyses, which we required to combine studies in our meta-analysis; instead, we report their unadjusted analyses for which there was an estimate of variance.^{7,9} The statistical significance of the findings of these 2 studies did not change between the unadjusted analyses.^{7,9}

Five of the 8 studies showed statistically significant higher payments for care at private for-profit hospitals,^{9,11-14} 1 showed statistically significant lower payments for care at private for-profit hospitals,⁸ 1 showed a nonsignificant trend toward higher payments for care at private for-profit hospitals⁷ and 1 showed a nonsignificant trend toward lower payments for care at private for-profit hospitals (Fig. 2).¹⁰ Funnel plots did not suggest publication bias.

Our primary meta-analysis demonstrated that private forprofit hospitals were associated with higher payments for care (relative payments for care 1.19, 95% confidence interval [CI] 1.07–1.33, p = 0.001). There was large heterogeneity

Table 1: Excluded studies initially thought to be potentially eligible

Problems that precluded study inclusion	Efforts to resolve problems		
Six studies evaluated payments for care in PFP and PNFP hospitals, but no measures of variance were available ¹⁵⁻²⁰	Five authors were unable to provide the data, ^{15,16,18-20} and we were unable to contact 1 author. ¹⁷ Two publications reported statistically significant higher payments for care in PFP hospitals, ^{15,16} 1 publication reported statistically significant higher payments for care in PFP hospitals for 11 of the 13 DRGs evaluated, ¹⁷ 1 publication reported a trend toward higher payments for care in PFP hospitals, ¹⁸ and 2 publications did not report statistical analyses but demonstrated higher payments for care in PFP hospitals ^{19,20}		
Three studies evaluated payments for care in PFP and NFP hospitals, but the NFP hospitals were a mixture of public and private NFP hospitals ²¹⁻²³	All 3 authors were contacted but were unable to provide the data to compare the PFP and PNFP hospitals. Two publications reported statistically significant higher payments for care in PFP hospitals, ^{21,22} and 1 publication did not report statistical findings but demonstrated higher payments for care in PFP hospitals ²³		
One study evaluated charges for care at PFP and PNFP hospitals ²⁴	The author was contacted but was unable to provide payment data. This study reported statistically significant higher charges for care in PFP hospitals ²⁴		

Note: PFP = private for-profit, PNFP = private not-for-profit, NFP = not-for-profit, DRG = diagnosis related group.

across the study results ($I^2 = 90\%$). The lone study, by Kauer, that showed statistically significant higher payments for care at private not-for-profit hospitals compared hospitals owned by not-for-profit organizations but run by a for-profit firm with hospitals owned and operated by private for-profit organizations.⁸ Because this study design is different than that of the other studies, we undertook a heterogeneity test that evaluated the difference between Kauer's estimate of payments for care (relative payments for care 0.93, 95% CI 0.88-0.99) and the estimate from the other studies (relative payments for care 1.24, 95% CI 1.11-1.39), and this test was statistically significant (p = 0.048). Even without Kauer's study, large heterogeneity across the study results (I² = 86%) persisted. Only 1 of our predefined hypotheses (specialty versus general hospitals) helped to explain some of the persistent heterogeneity (p = 0.02 for the difference between these subgroup summary estimates). Pooled estimates from both the 3 studies that evaluated specialty hospitals and the 5 studies that evaluated general hospitals showed higher payments for care at private for-profit hospitals (relative payments for care 1.48, 95% CI 1.15-1.89, and 1.11, 95% CI 1.00–1.23, respectively).

The 2 studies with the most extensive adjustment for potential confounders — the study by Sloan and associates¹² and the study by Keeler and associates¹³ — reported statistically significant higher payments for care at the private forprofit hospitals than at the private not-for-profit hospitals (relative payments for care 1.51, 95% CI 1.17–1.94, and 1.13, 95% CI 1.09–1.16, respectively).

Interpretation

We identified 8 observational studies that compared payments for care at private for-profit and private not-forprofit hospitals. These studies altogether involved more than 350 000 patients and included a median of 324 hospitals each. Five of the studies and our pooled analysis demonstrated statistically significant higher payments for care at private for-profit hospitals than at private not-forprofit hospitals.

Our systematic review has several strengths. We undertook a comprehensive search using 6 strategies to identify studies for our review, masked study results before determining study eligibility, conducted eligibility decisions and data abstraction in duplicate and demonstrated a high degree of agreement. For 7 of the 8 studies the investigators provided or confirmed data.

Our systematic review has several limitations. We did not identify any randomized controlled trials. It is unlikely that patients will ever be randomly assigned to private forprofit and private not-for-profit health care delivery systems. Therefore, the strongest realistic design for addressing our question is an observational study.

The main limitation of observational data is the potential for confounding. Six of the 8 observational studies adjusted for potential confounders. The statistical significance of the findings within the 2 studies for which we present the unadjusted analyses did not vary between the unadjusted and adjusted analyses.^{7,9} The 2 studies with the most extensive adjustment for potential confounders both reported statistically significant higher payments for care at the private for-profit hospitals (relative payments for care 1.51 and 1.13).^{12,13} For example, the study by Sloan and associates adjusted for age, sex, education, ethnicity, marital status, income, community living, number of activities of daily living, cognitive awareness, bladder/bowel control, comorbidity, primary diagnosis at index admission, market characteristics (population per square mile, Herfindahl index, Medicare hospital wage index, Health Maintenance Organization market share, hospital beds per 100 popula-

Table 2: Characteristics of studies included in the systematic review					
Study*	Type of hospital	Sources of payments	Date when data collection was initiated	Date when data collection was completed	Follow-up period for individual patients
Van Ness ⁷	General acute care	Public and private	01/01/1980	31/12/1981	In hospital
Kauer [®]	General acute care	Public and private	01/01/1981	31/12/1984	In hospital
Dickey ⁹	Mixed†	Private	01/07/1985	30/06/1987	In hospital
Dranove et al^{10}	General acute care	Public and private	Fiscal year 1983	Fiscal year 1992	In hospital
McCue et al^{11}	Psychiatric acute care	Public and private	Fiscal year 1986	Fiscal period ending 1990	In hospital
Sloan et al ¹²	General acute care	Medicare	01/01/1983	31/12/1994	6 mo
Keeler et al ¹³	General acute care	Medicaid and private	01/01/1986	31/12/1994	In hospital
McCue et al ¹⁴	Rehabilitation	Public and private	Fiscal year 1989	Fiscal year ending 1992	In hospital

*The studies are in chronological order by midpoint of the data collection period.

†Mixed = general, major teaching, free-standing psychiatric or free-standing substance abuse.

Study*	Sampling method	Data source	Case mix	Factors controlled for in the analysis
Van Ness ⁷	All acute care general hospitals that reported annual utilization and financial data to the California HFC. Excluded hospitals: psychiatric, children's, university teaching, state, government, Kaiser Foundation, dental, and hospitals managed under contract by for-profit or not-for-profit chains	California HFC Commission, HCFA	No restriction	Unadjusted
Kauer®	Hospitals randomly selected from all domestic owned or contract managed HCA hospitals that were acute care medical/surgical hospitals. Excluded hospitals: those with missing critical data, specialty hospitals, flagship hospitals and hospitals that were with the HCA system for < 3 yr	Audited year-end financial statements, HCA operating indicators reports, and Federal Register	No restriction	Case mix, no. of beds, occupancy, wage index, ancillary services, year
Dickey ⁹	Psychiatric and substance abuse admissions of employees and their dependents < 65 yr old from 2 large national corporations with generous indemnity health plans that included nondiscriminatory unlimited inpatient mental health benefits	Paid claims data from health plans	Psychiatric and substance abuse disorders	Unadjusted
Dranove et al ¹⁰	Private short-term hospitals in California that reported data to the California OSHPD and that had enough Medicaid patients to allow for reliable measures of service levels during fiscal years 1983 and 1992	Hospital disclosure files and discharge data files of the California OSHPD	No restriction	Case mix
McCue et al ¹¹	Short-term psychiatric hospitals reported by the national association of Private Psychiatric Hospitals in 1989; matched by profit status according to location in same county, standard metropolitan statistical area or wage index	HCFA data	Psychiatric disorders	Matching process was validated by testing the means of pair differences for a set of market measures, including wage index, county population, no. of psychiatric beds in county, age population categories and total no. of beds
Sloan et al ¹²	All Medicare patients with 1 of 4 diagnoses admitted to a nonfederal, general hospital with LOS < 92 d who were also included in Long-Term Care Survey (voluntary national Medicare survey of patients > 65 yr old with ≥ 1 limitation of ADL or instrumental ADL, undertaken in 1982, 1984, 1989 and 1994)	Medicare claims data were merged with National Long-Term Care Survey data	Hip fracture, stroke, coronary artery disease, congestive heart failure	Age, sex, education, ethnicity, marital status, income, community living, no. of ADLs, cognitive awareness, bladder/bowel control, comorbidity, primary diagnosis at index admission, market characteristics (population per square mile, Herfindahl index, Medicare hospital wage index, HMO market share, hospital beds per 100 population), year of index admission, no. of hospital beds and teaching status
Keeler et al ¹³	Non-Medicare patients admitted with 1 of 10 common medical problems to a California hospital that submitted discharge data to the California OSHPD during 1986, 1989, 1992 and 1994. Excluded hospitals: Kaiser Permanente, military, specialty, psychiatric, rehabilitation and long-term care	Annual uniform discharge data and hospital disclosure data from the California OSHPD	Patients admitted with 1 of 10 common medical problems (e.g., cerebrovascular disease, pneumonia, heart failure)	Patient characteristics, case mix, LOS, percentage of admissions in each hospital covered by Medicare/Medicaid, capital ratio (total assets / total operating expenses), teaching status, county level measures (e.g., population per square mile, per capita income in 1988), Medicare prospective payment system wage price index, Herfindahl- Hirschman Index, year and no. of admissions
McCue et al ¹⁴	HCFA Minimum Cost Data Set was used to select 2 sample groups (existing and new rehabilitation hospitals), which differed in the establishment of their target reimbursement level under TEFRA	HCFA data	No restriction	Case mix

Table 3: Methodology of studies included in the systematic review

Note: HFC = Health Facilities Commission, HCFA = Health Care Financing Administration, HCA = Hospital Corporation of America, a private for-profit company, OSHPD = Office of Statewide Health Planning and Development, LOS = length of stay, ADL = activity of daily living, HMO = Health Maintenance Organization, TEFRA = Tax Equity and Fiscal Responsibility Act. *The studies are in chronological order by midpoint of the data collection period.

tion), year of index admission, number of hospital beds and hospital teaching status. $^{\rm 12}$

Another limitation of our systematic review was that we were unable to include 6 studies because the investigators failed to provide an estimate of variance, 3 studies because the investigators either no longer had the data or could not rerun their analyses to remove the public notfor-profit hospitals, and 1 study because the investigators had only charges and not payment data. All 10 of these studies, however, showed higher payments or charges for care at private for-profit hospitals than at not-for-profit hospitals, and in 6 of the 10, the differences were statistically significant.

Our pooled analysis showed significant variability in the direction and magnitude of effect among the studies. It is common practice to pool results with significant heterogeneity, but one may question the advisability of doing so. Although the inference that private for-profit hospitals result in higher payments for care is secure, heterogeneity in results suggests that the magnitude of the effect may differ according to circumstances. In the presence of unexplained heterogeneity, inferences associated with pooled estimates are weaker; nonetheless, these estimates provide the best available estimate of the average effect.²⁵ We present these results as we believe they constitute useful information for decision-makers.

Our pooled analysis and 5 of the studies included in our

systematic review demonstrated statistically significant higher payments for care at private for-profit hospitals than at private not-for-profit hospitals, whereas only 1 study demonstrated the opposite.⁸ These results are completely consistent with those of studies that we were unable to include because of technical reasons (Table 1). The lone study demonstrating lower payments for care at private forprofit hospitals compared them with hospitals owned by private not-for-profit organizations but run by a private for-profit firm.⁸ In essence, this is a comparison between different modes of for-profit management.

Why would private for-profit hospitals have higher payments for care than private not-for-profit hospitals? One potential explanation could be that they are providing superior care. However, our previous meta-analysis involving over 38 million patients demonstrated that private for-profit hospitals have higher risk-adjusted mortality rates.² Our meta-analysis involving over 500 000 patients receiving hemodialysis also revealed higher risk-adjusted mortality rates at private for-profit dialysis facilities.²⁶

The likely explanation is the necessity to generate revenues to satisfy investors, a requirement absent in private not-for-profit hospitals. Private for-profit hospitals are also burdened with a 6% absolute increase in the proportion of hospital spending devoted to administration as compared with private not-for-profit hospitals.²⁷ Further, executive

Study*	No. of facilities	No. of patients	% weight		PFP/PNFP payment ratio (95% CI)
Van Ness ⁷	333	NA	13.7	⊢ ← -1	1.09 (0.98–1.22)
Kauer [®]	56	NA	15.1	⊢◆┤	0.93 (0.88-0.99)
Dickey ⁹	342	561	8.9	⊢ ◆	1.73 (1.36–2.20)
Dranove et al ¹⁰	314	NA	14.4	⊢┥	0.98 (0.90–1.07)
McCue et al ¹¹	84	NA	10.5		1.62 (1.34–1.97)
Sloan et al ¹²	2 360†	7 079	8.4		1.51 (1.17–1.94)
Keeler et al ¹³	358†	384 000	15.8	 ♦	1.13 (1.09–1.16)
McCue et al ¹⁴	131	NA	13.2		1.20 (1.06–1.36)
Pooled random eff $I^2 = 0.903$	ects estimate (<i>p</i> = 0.001)			1.19 (1.07–1.33)
	ـــ 0.33		0.67	1.0 1.33 1.67 2.0	2.33 2.67 3.0
		Lo	wer payments PFP hospitals	Higher paymer at PFP hospita	nts

Fig. 2: Relative payments for care at private for-profit (PFP) and private not-for-profit (PNFP) hospitals. Note: CI = confidence interval. *The studies are in chronological order by midpoint of the data collection period. †Approximation from investigator.

bonus incentives are over 20% higher at private for-profit than at private not-for-profit hospitals.²⁸

We offer 2 reasons why our results may actually underestimate the association between private for-profit hospitals and higher payments for care. First, all but 2 of the studies in our systematic review adjusted for case mix. Therefore, our results do not capture any increase in payments for care resulting from inappropriate upcoding of patient diagnoses to enhance reimbursement. Private for-profit hospitals manifest higher upcoding of patient diagnoses than do private not-for-profit hospitals.²⁹

Second, the studies in our systematic review did not explicitly address issues of fraud (e.g., performance of unnecessary surgeries, billing for services not provided, inappropriate detainment of psychiatric patients for billing purposes),³⁰ which can increase both the direct costs of care and the indirect costs related to investigating and prosecuting offenders. The multimillion-dollar fraud lawsuits in the United States have overwhelmingly been against private for-profit hospitals.^{31,32} It is likely, therefore, that we are underestimating the true association between private for-profit hospitals and higher payments for care.

Many countries, like Canada, are debating choices about private for-profit and private not-for-profit health care delivery. How important is a relative increase in payment for care of 19%? Canada currently spends \$120 billion annually on health care, and hospital care accounts for 32% of overall expenditures.³³ If we were to convert half of our hospitals to private for-profit institutions, our results suggest that we would pay approximately an extra \$3.6 billion annually.

Given the differences in the structure of Canadian and US health care systems, one might question the applicability of our results to Canada. The structure of US health care has, however, changed significantly over time (e.g., the introduction of prospective payment systems for Medicare patients and managed health care). The 5 studies that demonstrated statistically significant higher payments for care at private for-profit hospitals included data from 1983 through to the end of 1994, and thus included results from both before and after these changes. Furthermore, these 5 studies had variations in their sources of payments (i.e., Medicare, private insurance or both). These findings suggest that the higher payments for care at private for-profit hospitals are manifest within a variety of health care contexts. Finally, should Canada open the door to private forprofit hospitals, the very same large US hospital chains that have generated the data included in this systematic review will soon be purchasing Canadian private for-profit hospitals. In summary, it is likely that our results are generalizable to the Canadian context.

For-profit hospitals result in both higher mortality rates and greater payments for care than do not-for-profit hospitals. The evidence strongly supports a policy of not-forprofit health care delivery at the hospital level.

This article has been peer reviewed.

From the Department of Medicine (Devereaux, Cook, Guyatt), the Department of Clinical Epidemiology and Biostatistics (Devereaux, Heels-Ansdell, Lacchetti, Haines, Cook, Walter, Bhandari, Lavis, Stoddart, Guyatt), the Faculty of Medicine (Stone), the Department of Surgery (Bhandari), the Department of Political Science (Lavis) and the Centre for Health Economics and Policy Analysis (Stoddart), McMaster University, Hamilton, Ont; the Faculty of Medicine (Ravindran), Inner City Health Research Unit, St. Michael's Hospital, and the Department of Medicine (Bayoumi), Cancer Care Ontario and the Department of Health Policy Management and Evaluation, University of Toronto (Sullivan), and AXON Clinical Research (McDonald), Toronto, Ont.; the Department of Medicine, University of Western Ontario, London, Ont. (Burns); the Department of Medicine, University of Ottawa, Ottawa, Ont. (Patel); the Departments of Medicine and Social and Preventive Medicine, University at Buffalo, Buffalo, NY (Schünemann); the Vancouver Coastal Health Research Institute and the Department of Anesthesia, University of British Columbia, Vancouver, BC (Choi)

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Correspondence to: Dr. P.J. Devereaux, McMaster University, Faculty of Health Sciences, Department of Clinical Epidemiology and Biostatistics, Rm. 2C8, 1200 Main St. W, Hamilton ON L8N 3Z5; fax 905 526-1353; philipj@mcmaster.ca

